

HEDGING OF OPTIONS UNDER MEAN-SQUARE CRITERION AND SEMI-MARKOV VOLATILITY

ХЕДЖУВАННЯ ОПЦІОНУ ЗА УМОВ СЕРЕДНЬОКВАДРАТИЧНОГО КРИТЕРІЮ ТА ПІВМАРКОВСЬКИХ МІНЛИВОСТЕЙ

We consider a problem of hedging of the European call option for a model such that appreciation rate and volatility are functions of a semi-Markov process. In such a model, the market is incomplete.

Розглядається задача хеджування Європейського опціону купівлі для моделі з нормою повернення та коефіцієнтом мінливості, що залежать від півмарковського процесу. В такій моделі ринок є неповним.

1. Introduction. In famous Black – Scholes model, which is used for evaluation of option prices, it is supposed that the dynamic of stocks prices is set by the linear stochastic differential equation

$$dS_t = aS_t dt + \sigma S_t dW_t,$$

where a and σ are deterministic functions (in the simplest case – constants).

We suppose that, in our model, the coefficients a and σ , appreciation rate and volatility respectively, are depended on a semi-Markov process X_t , which doesn't depend on standard Wiener process W_t . We consider the hedging problem of the European call option with terminal payment $H = f(S_T)$.

Since the additional source of randomness exists (the semi-Markov process X_t) in addition to the Wiener process W_t , the market is incomplete and perfect hedging is not possible. We find a strategy, which locally minimizes the risk.

2. Description of the model and preliminary notions. Let $(\Omega, \mathcal{F}, \mathbf{F} = (\mathcal{F}_t), \mathcal{P})$ be a probability space. We suppose that risk assets (stock) exist and their price evolution is given by the following stochastic differential equation [5]:

$$dS_t = a(X_t)S_t dt + \sigma(X_t)S_t dW_t, \quad (1)$$

where W_t is a Wiener process, X_t is some observed variable, which is described by a semi-Markov process [4, 5] with the phase space $(\mathbf{X}, \mathcal{X})$, $X_t := X_{\nu(t)}$, $\nu(t) := \max \{n : \tau_n \leq t\}$, $\tau_n := \sum_{k=1}^n \theta_k$, $(X_n, \theta_n; n \geq 0)$ is a Markov renewal process, $\mathcal{P}\{\omega : X_{n+1} \in A, \theta_{n+1} \leq t / X_n = x\} = P(x, A) \cdot G_x(t)$, $x \in \mathbf{X}$, $A \in \mathcal{X}$, $t \geq 0$. We suppose that $G_x(t)$ is a differentiable function of t and $g_x(t) := dG_x(t)/dt$, $\forall x \in \mathbf{X}$. Coefficients a and σ are measurable functions on \mathbf{X} , $\sigma > 0$, processes W_t and X_t are independent, and filtration \mathbf{F} is generated by X_t and W_t .

We solve the problem of hedging of the European call option which is sold at the moment $t = 0$, with terminal payment $H = f(S_T)$ at a cancellation moment T . It's considered that $EH^{2+\varepsilon} < +\infty$, $\varepsilon > 0$.

Besides the risk assets, we have nonrisk assets (bond or bank account), and we suppose that its price is a constant and is equal to one (without of loss generality) at all moments (i.e., percentage rate is equal to zero).

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Stock exchange strategy (SES) π is a pair (γ, β) , where $\gamma = (\gamma_t)$ is such a predictable process that

$$E \int_0^T \gamma_t^2 \sigma^2(X_t) S_t^2 dt + E \left(\int_0^T |\gamma_t| |a(X_t)| S_t dt \right)^2 < +\infty, \quad (2)$$

$\beta = (\beta_t)$ is a coordinated process, $E \beta_t^2 < +\infty, \forall t \leq T$.

SES defines a portfolio with a number of units of risk assets γ_t (which the holder has at the moment t) and with a number of means that were invested in bonds at the moment t .

The value process $V(\pi)$ of the portfolio with respect to the strategy π is defined as

$$V_t(\pi) = \gamma_t S_t + \beta_t, \quad (3)$$

and the cost process

$$C_t(\pi) = V_t(\pi) - \int_0^t \gamma_r dS_r, \quad (4)$$

SES π is said to be H -admissible if $V_T(\pi) = H$; SES π is said to be self-financing (or mean-value self-financing) if the cost process $C_t(\pi)$ is a constantine time (or martingale).

The residual risk is defined by the formula

$$\mathcal{R}_t(\pi) := E \{ |C_T(\pi) - C_t(\pi)|^2 / \mathcal{F}_t \}. \quad (5)$$

The SES H -admissible strategy π^* is called a risk-minimizing if, for any H -admissible SES π and for any t ,

$$\mathcal{R}_t(\pi^*) \leq \mathcal{R}_t(\pi)$$

The purpose of this article is to find a risk-minimizing H -admissible SES.

It was stated in [2] that existence of a risk-minimizing H -admissible SES is equivalent to the existence of the expansion of the terminal payment H in the form

$$H = H_0 + \int_0^T \gamma_r^H dS_r + L_T^H \quad (\mathcal{P}\text{-a.s.}), \quad (6)$$

where $H_0 \in L^2(\mathcal{F}_0, \mathcal{P})$, γ_{Ht}^H satisfies (2), and L_T^H is a square-integrable martingale, which is orthogonal to the martingale component S . Then the γ -component of a risk-minimizing strategy π is $\gamma = \gamma^H$ and $C_\pi = H_0 + L^H$.

To obtain expansion (6), it is introduce a minimal martingale measure $\tilde{\mathcal{P}}$ [2].

For our model (1), the minimal martingale measure is $\tilde{\mathcal{P}} = \rho_T \mathcal{P}$, where the density ρ_T is defined by the equality

$$\rho_T = \exp \left\{ - \int_0^T \frac{a(X_r)}{\sigma(X_r)} dW_r - \frac{1}{2} \int_0^T \frac{a^2(X_r)}{\sigma^2(X_r)} dr \right\}. \quad (7)$$

In our situation, when the process S_t has continuous paths, the process $a(X)/\sigma(X)$ is bounded and $H \in L^{2+\varepsilon}(\mathcal{P})$, $\varepsilon > 0$, and the desired expansion (6) can

be obtained from Kunita – Watanabe's expansion (as $t = T$) with respect to the measure \tilde{P} :

$$\tilde{E}(H | \mathcal{F}_t) = \tilde{E}H + \int_0^t \tilde{\gamma}_r^H dS_r + \tilde{L}_t^H. \quad (8)$$

In such a way, it is need to find the Kunita – Watanabe's expansion. Let's introduce some notations.

A jump measure for X_t has the following form:

$$\mu([0, t] \times A) = \sum_{n \geq 0} I(X_n \in A, \tau_n \leq t), \quad A \in \mathcal{X}, \quad t \geq 0. \quad (9)$$

It's known [4], that \mathbf{F} -dual predictable projection for μ has the form:

$$\nu(dt, dy) = \sum_{n \geq 0} I(\tau_n < t \leq \tau_{n+1}) \frac{P(X_n, dy) g_{X_n}(t)}{\bar{G}_{X_n}(t)} dt, \quad (10)$$

where $\bar{G}_x(t) := 1 - G_x(t)$, $g_x(t) := dG_x(t)/dt$, $\forall x \in \mathbf{X}$, $t \geq 0$.

For given $H \in L_2(\tilde{P})$, we find the expansion

$$\tilde{E}(H | \mathcal{F}_t) = \tilde{E}H + \int_0^t \tilde{\gamma}_r^H dS_r + \int_0^t \int_{\mathbf{X}} \tilde{\Psi}^H(r, y) (\mu - \nu)(ds, dy). \quad (11)$$

We note that the last integral in (11) is \tilde{P} -orthogonal to \mathcal{S} (it's an (\mathbf{F}, \tilde{P}) -martingale), and the uniqueness of the Kunita – Watanabe's expansion is a guarantee that (11) is the desired expansion of (8) for our case.

Finally, a risk-minimizing H -admissible strategy π^* is defined by $\pi^* = (\gamma^*, \beta^*)$, where $\gamma^* = \tilde{\gamma}^H$, and β^* is such that $V_t(\pi^*) = \tilde{E}(H | \mathcal{F}_t)$, i.e.,

$$\beta_t^* = \tilde{E}(H | \mathcal{F}_t) - \tilde{\gamma}_t^H S_t. \quad (12)$$

In the next section we will obtain an exact representation for $\pi^* = (\gamma^*, \beta^*)$.

3. The result. Let $f(z)$ be a function such that $|f(z)| \leq c \cdot (1 + |z|)^m$ for some $m \geq 0$. Let's consider a function $u(t, z, x)$ on $[0, T] \times \mathbf{R}_+ \times \mathbf{X}$ such that it is a solution of the Cauchy problem:

$$\begin{cases} u_t(t, z, x) + \frac{1}{2} \sigma^2(x) \cdot z^2 \cdot u_{zz}(t, z, x) + Au(t, z, x) = 0, \\ u(T, z, x) = f(z). \end{cases} \quad (13)$$

where

$$Au(t, z, x) := \frac{g_x(t)}{\bar{G}_x(t)} \int_{\mathbf{X}} P(x, dy) [u(t, z, y) - u(t, z, x)]. \quad (14)$$

Theorem 1. The risk-minimizing H -admissible stock exchange strategy $\pi^* = (\gamma^*, \beta^*)$ is given by the following formula:

$$\gamma_t^* = u_z(t, S_t, X_t), \quad (15)$$

$$\beta_t^* = V_t(\pi^*) - \gamma_t^* \cdot S_t,$$

where

$$V_t(\pi^*) = \tilde{E} f(S_T) + \int_0^t u_-(r, S_r, X_r) dS_r + \int_0^t \int_{\mathbf{X}} \psi(r, y) (\mu - \nu)(dr, dy), \quad (16)$$

$$\psi(r, y) = u(r, S_r, y) - u(r, S_r, X_{r-}).$$

The residual risk process has the form

$$\mathcal{R}_t(\pi^*) = E \left(\int_t^T [A u^2(r, S_r, X_r) - 2u(r, S_r, X_r) A u(r, S_r, X_r)] ds / \mathcal{F}_t \right).$$

In particular, the residual risk at the moment $t=0$ is equal to

$$\mathcal{R}_0(\pi^*) = E \left(\int_0^T [A u^2(r, S_r, X_r) - 2u(r, S_r, X_r) A u(r, S_r, X_r)] ds \right), \quad (17)$$

where the operator A was defined in (14).

4. Proof. By applying Ito's formula to the solution of (13) we obtain:

$$f(S_T) = u(T, S_T, X_T) = u(0, z, x) + \int_0^T u_-(r, S_r, X_r) dS_r +$$

$$+ \int_0^T \left[u_t(r, S_r, X_r) + \frac{1}{2} \sigma^2(X_r) (S_r)^2 u_{zz}(r, S_r, X_r) \right] dr +$$

$$+ \sum_{r \leq T} [u(r, S_r, X_r) - u(r-, S_{r-}, X_{r-})]. \quad (18)$$

We note that, for any function h on $[0, T] \times \mathbf{X}$, right-continuous and left-limit of t , we have

$$\sum_{r \leq T} [h(r, X_r) - h(r-, X_{r-})] = \int_0^T \int_{\mathbf{X}} [h(r, y) - h(r-, X_{r-})] \mu(dr, dy) =$$

$$= \int_0^T \int_{\mathbf{X}} [h(r, y) - h(r-, X_{r-})] (\mu - \nu)(dr, dy) +$$

$$+ \int_0^T \int_{\mathbf{X}} [h(r, y) - h(r-, X_{r-})] \nu(dr, dy) =$$

$$= \int_0^T \int_{\mathbf{X}} [h(r, y) - h(r-, X_{r-})] (\mu - \nu)(dr, dy) +$$

$$+ \int_0^T \int_{\mathbf{X}} \frac{g_{X_r}(r)}{G_{X_r}(r)} P(X_{r-}, dy) [h(r, y) - h(r-, X_{r-})] dr =$$

$$= \int_0^T \int_{\mathbf{X}} [h(r, y) - h(r-, X_{r-})] (\mu - \nu)(dr, dy) + \int_0^T A h(r-, X_{r-}) dr \quad (19)$$

(see (10) and (14)).

Hence, from (13), (18), and (19) we obtain

$$\begin{aligned}
 f(S_T) &= u(T, S_T, X_T) = u(0, z, v) + \int_0^T u_z(r, S_r, X_r) dS_r + \\
 &+ \int_0^T \int_{\mathbf{X}} [u(r, S_r, X_r) - u(r, S_r, X_{r-})](\mu - \nu)(dr, dy), \quad (20)
 \end{aligned}$$

and relations (15), (16) are valid.

Residual risk process can be expressed in the following way (see (11)):

$$\begin{aligned}
 \mathcal{R}_t(\pi^*) &= E \left(\left[\int_t^T \int_{\mathbf{X}} \psi(r, y)(\mu - \nu)(dr, dy) \right]^2 / \mathcal{F}_t \right) = \\
 &= E \left(\int_t^T \int_{\mathbf{X}} [u(r, S_r, y) - u(r, S_r, X_r)]^2 \nu(dr, dy) / \mathcal{F}_t \right) = \\
 &= E \left(\int_t^T \int_{\mathbf{X}} \frac{g_{X_r}(r)}{G_{X_r}(r)} P(X_{r-}, dy) [u(r, S_r, y) - u(r, S_r, X_{r-})]^2 / \mathcal{F}_t \right) = \\
 &= E \left(\int_t^T [A u^2(r, S_r, X_r) - 2u(r, S_r, X_r) A u(r, S_r, X_r)] dr / \mathcal{F}_t \right).
 \end{aligned}$$

and the theorem is proved.

We must only prove that the solution of Cauchy problem (13) exists. We do it in the next section.

5. Random evolution approach. Let \tilde{S}_t be a solution of the stochastic differential equation

$$d\tilde{S}_t = \sigma(X_t)\tilde{S}_t dW_t, \quad \tilde{S}_0 = z. \quad (21)$$

This solution has the form

$$\tilde{S}_t = z \exp \left\{ \int_0^t \sigma(X_s) dW_s - \frac{1}{2} \int_0^t \sigma^2(X_s) ds \right\}. \quad (22)$$

We note that \tilde{S}_t is a continuous semi-Markov random evolution [5: 4, p. 77] $V^{\tilde{S}}(t)$:

$$V^{\tilde{S}}(t)f(z) := E[f(\tilde{S}_t) / X(s), 0 \leq s \leq t], \quad (23)$$

i.e., random evolution underlying the semi-Markov process X_t . This evolution is generated by the following generating operators:

$$\Gamma(y)f(z) := \frac{1}{2} \sigma^2(y) z^2 f''(z) \quad \forall f(z) \in \mathbf{C}^2(\mathbf{R}_+). \quad (24)$$

Further, let's consider the following process $(X_t, t - \tau_{\text{exit}})$. It is a Markov process on $\mathbf{X} \times \mathbf{R}_+$ with the infinitesimal operator $\tilde{A} = A + d/dt$, where A is defined in (14).

The expectation for the random evolution $V^{\tilde{S}}(t)$ of Markov process $(X_t, t - \tau_{\text{exit}})$ satisfies the following equation:

$$v(t, z, x) := E[V^{\tilde{S}}(t)f(z, X_t, t - \tau_{\text{exit}})]$$

$$\begin{cases} \frac{dv}{dt} = \Gamma(x)v + \tilde{A}v, \\ v(0, z, x) = f(z, x, 0), \end{cases} \quad (25)$$

where $\Gamma(x)$ is defined in (24).

Let $(\tilde{S}_t^{\pm, x}, X_t^x, t - \tau_{x(t)})$ be a Markov process with the initial point $(z, x, 0)$ with the first component \tilde{S}_t^{\pm} in (21). X_t^x be a semi-Markov process. From (21)–(25) we obtain the following result:

Lemma 1. *Function*

$$u(t, z, x) := E f(\tilde{S}_t^{\pm, x}) \quad (26)$$

is a solution of problem (13).

Proof. From (22) we have

$$u(T-t, z, x) = E f(\tilde{S}_t^{\pm, x}) = \int f(y) y^{-1} h(y; t, z, x) dy, \quad (27)$$

where

$$h(y; t, z, x) = \int \varphi\left(\xi, \ln \frac{y}{z} + \frac{1}{2}\xi\right) F_t^x(d\xi) = E\varphi\left(z_t^x, \ln \frac{y}{z} + \frac{1}{2}z_t^x\right), \quad (28)$$

$\varphi(t, x) = (2\pi t)^{-1/2} \exp\{-x^2/2t\}$. F_t^x is a distribution of the random variable

$$Z_t^x = \int_0^t \sigma^2(X_r^x) dr. \quad (29)$$

Let v be a solution of the equation

$$v_z(t, \eta, x) = \sigma^2(x)v_z(t, \eta, x) + \tilde{A}v(t, \eta, x) \quad (30)$$

with the initial condition $v(0, \eta, x) = g(\eta)$. From Ito's formula we have that

$$v(t, \eta, x) = E g(\eta + Z_t^x). \quad (31)$$

By substituting in this formula $g(\eta) = \varphi(\eta - \ln(y/z), (\ln(y/z) + \eta)/2)$, we obtain that

$$h(y; t, z, x) = v\left(\ln \frac{y}{z}, t, x\right). \quad (32)$$

From (30) and (32) we have:

$$\begin{aligned} h_t(y; t, z, x) &= \sigma^2(x) E \left[\varphi_t\left(Z_t^x, \ln \frac{y}{z} + \frac{1}{2}Z_t^x\right) + \right. \\ &\quad \left. + \frac{1}{2} \varphi_z\left(Z_t^x, \ln \frac{y}{z} + \frac{1}{2}Z_t^x\right) \right] + \tilde{A}h(y; t, z, x), \end{aligned}$$

Differentiation of (28) gives the equality

$$h_{zz} = z^{-2} E[\varphi_{zz} + \varphi_z].$$

Since φ satisfies the heat equation $\varphi_t = 1/2 \varphi_{zz}$,

$$h_t = \frac{1}{2} \sigma^2(x) z^2 h_{zz} + \tilde{A}h.$$

Hence, the function u in (26) is a solution of (13).

Remark 1. Let's define the following process:

$$m_t := f(\tilde{S}_t, X_t, t - \tau_{v(t)}) - f(z, x, 0) - \int_0^t \left(A + \frac{d}{dr} \right) f(\tilde{S}_r, X_r, r - \tau_{v(r)}) dr. \quad (33)$$

It's an \mathcal{F}_t -martingale, where $\mathcal{F}_t := \sigma\{X_s, W_s; 0 \leq s \leq t\}$. It's quadratic variation is equal to

$$\begin{aligned} \langle m_t \rangle &= \int_0^t \left[\left(A + \frac{d}{dr} \right) f^2(\tilde{S}_r, X_r, r - \tau_{v(r)}) - \right. \\ &\quad \left. - 2f(\tilde{S}_r, X_r, r - \tau_{v(r)}) \left(A + \frac{d}{dr} \right) f(\tilde{S}_r, X_r, r - \tau_{v(r)}) \right] dr = \\ &= \int_0^t [Af^2(\tilde{S}_r, X_r, r - \tau_{v(r)}) - 2f(\tilde{S}_r, X_r, r - \tau_{v(r)})Af(\tilde{S}_r, X_r, r - \tau_{v(r)})] dr. \quad (34) \end{aligned}$$

In such a way, from (17) and (34), it follows that $\mathcal{R}_{\mathcal{G}}(\pi^*) = \langle m_T \rangle$ with the function u replacing f in (33).

Remark 2. In the Markov case, the operator A in (14) has the following form:

$$Af(x) = \lambda(x) \int_{\mathbf{X}} P(x, dy) [f(y) - f(x)], \quad (35)$$

where $\lambda(x)$ are an intensities of jumps of the jump Markov process X_t ; in this case

$$Q(x, A, t) = P(x, dy)(1 - e^{-\lambda(x)t}),$$

$$G_x(t) = 1 - e^{-\lambda(x)t},$$

$$\bar{G}_x(t) = e^{-\lambda(x)t},$$

$$g_x(t) = \lambda(x)e^{-\lambda(x)t},$$

and

$$\frac{g_x(t)}{\bar{G}_x(t)} = \lambda(x).$$

In this way, the operator A in (35) is an infinitesimal operator of the jump Markov process X_t .

Corollary 1. Initial capital for hedging strategy in our model is defined by the formula:

$$V_0(\pi) = \bar{E} f(S_T) = \int \left(\int f(y) y^{-1} \varphi \left(\eta, \ln \frac{y}{z} + \frac{1}{2} \eta \right) dy \right) F_T^s(d\eta).$$

In particular, for the European call options $f(y) = (y - K)^+$, we have

$$V_0(\pi) = \int C_{BS}((z/T)^{1/2}, T) F_T^s(dz),$$

where $C_{BS}(\hat{\sigma}, T)$ is the Black - Scholes price for call option with volatility $\hat{\sigma}$.

i.e., $C_{BS}((z/T)^{1/2}, T) = S_0 \Phi(d_+) - K \Phi(d_-)$, where

$$d_{\pm} = \left[\ln \frac{S_0}{K} \pm \frac{S_0}{2} \right] / \sqrt{S_0}.$$

Corollary 2. Let $X = (1, 2)$, and $\nu(t)$ be a counting process for X_t , then

$$Z_T^1 = \int_0^T [\sigma^2(1)I(X_t=1) + \sigma^2(2)I(X_t=2)] dt = aT + bJ_T,$$

where

$$J_T = \int_0^T (-1)^{\nu(t)} dt,$$

$$a = \frac{1}{2} (\sigma^2(1) + \sigma^2(2)),$$

$$b = \frac{1}{2} (\sigma^2(1) - \sigma^2(2)).$$

1. Föllmer H., Sondermann D. Hedging of nonredundant contingent claims // Contributions to Mathematical Economics / Ed. W. Hildenbrand, A. MasColell. – Amsterdam–New York: North-Holland, 1986. – P. 205–223.
2. Föllmer H., Schweizer M. Hedging of contingent claims under incomplete information // Appl. Stoch. Anal. – New York–London: Gordon and Breach, 1991. – P. 389–414.
3. Di-Masi G., Kavanov Yu., Runggaldier W. Hedging of options under mean-square criterion and Markov volatility // Theory Probab. Appl. – 1994, – 39, N° 1, – P. 211–222.
4. Koroljuk V. S., Swishchuk A. V. Semi-Markov random evolutions. – Kiev: Nauk. Dumka, 1992. – 256 p.
5. Swishchuk A. V. Limit theorems for stochastic differential equations with semi-Markov switchings // Proceed. III Int. conf. "Evolution Stochastic Systems: Theory and Applications", Katsively, Ukraine. – 1992. – Utrecht–Moscow: VSP–TVP, 1994. – 14 p.

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